



Short Communication

# On the characteristic and moment generating functions of type-2 (Fréchet) and type-3 (reversed Weibull) distributions

G. Muraleedharan\* and C. Guedes Soares

Centre for Marine Technology and Ocean Engineering (Centec), Instituto Superior Técnico, Universidade de Lisboa, Av. Rovisco Pais, 1049-001, Lisbon, Portugal  
g.muraleedharan@centec.tenico.ulisboa.pt

Available online at: [www.isca.in](http://www.isca.in), [www.isca.me](http://www.isca.me)

Received 21<sup>st</sup> March 2017, revised 3<sup>rd</sup> April 2017, accepted 10<sup>th</sup> April 2017

## Abstract

We are able to derive for the first time the simplest forms of characteristic functions (CHF) and moment generating functions (MGF) of type-2 (Fréchet) and type-3 (reversed Weibull) in explicit closed forms by direct and unique methodology. CHF and MGF have wide applications in statistical theories such as in inversion and convolution.

**Keywords:** Characteristic Function, Moment generating function, GEV distribution, Fréchet distribution, reversed Weibull distribution.

## Introduction

The generalised extreme value distribution was introduced by Jenkinson<sup>1,2</sup> and recommended by Natural Environment Research Council of Great Britain<sup>3</sup>. The GEV distribution is the most widely accepted distribution for flood frequency data<sup>4-8</sup>.

Let  $X$  be a generalised extreme value (GEV) random variable represented by the probability density function as

$$f_X(x)dx = \frac{1}{\alpha} \left[ 1 - \kappa \left( \frac{x-\xi}{\alpha} \right) \right]^{\frac{1}{\kappa}-1} \exp \left\{ - \left[ 1 - \frac{\kappa(x-\xi)}{\alpha} \right]^{\frac{1}{\kappa}} \right\} dx \quad (1)$$

for range of  $x$ :  $-\infty < x \leq \xi + \frac{\alpha}{\kappa}$  if  $\kappa > 0$ ;  $-\infty < x < \infty$  if  $\kappa = 0$ ;  $\xi + \frac{\alpha}{\kappa} \leq x \leq \infty$  if  $\kappa < 0$

In this paper we derived a closed form for characteristic function (CHF) and moment generating function (MGF) of generalised extreme value distribution (GEV) for  $\kappa > 0$  (also known as type 3 EVD or reversed Weibull distribution) and type-2 EVD ( $\kappa < 0$ ; also known as Fréchet distribution).

When  $\kappa = 0$ , then

$$f_X(x)dx = \frac{1}{\alpha} \exp \left[ - \left( \frac{x-\xi}{\alpha} \right) \right] \exp \left\{ - \left[ \exp - \left( \frac{x-\xi}{\alpha} \right) \right] \right\} \quad (2)$$

is the Gumbel probability density function.

The probability density function of Fréchet distribution is given as:

$$f_X(x)dx = \frac{\kappa}{\alpha} \left( \frac{x-\xi}{\alpha} \right)^{-\kappa-1} \exp \left[ - \left( \frac{x-\xi}{\alpha} \right)^{-\kappa} \right] dx; x > \xi \quad (3)$$

$\xi$ ,  $\alpha$  and  $\kappa$  are respectively location, scale and shape parameters.

## Derivations of type-2 and type-3 extreme value distributions

**CHF of type-3 EVD:** The probability density function of generalised extreme value distribution (GEV) can be given as:

$$f_Z(z)dz = (1 - \kappa z)^{\frac{1}{\kappa}-1} \exp \left[ - (1 - \kappa z)^{\frac{1}{\kappa}} \right] dz, z = \frac{x-\xi}{\alpha}; -\infty < z \leq \frac{1}{\kappa}; \kappa > 0 \quad (4)$$

When  $\kappa > 0$ , then (1) is the probability density function of type-3 extreme value distribution and also known as the reversed Weibull distribution.

The CHF of (1) can be computed as

$$\phi_Z(t; \xi, \alpha, \kappa) = \int_R [\cos(taz)f(z) + i \sin(taz)f(z)] dz \quad (5)$$

$t$ -any arbitrary real constant

$$i = \sqrt{-1}$$

Integrating Ist term by parts  $\Rightarrow$

$$\int_R \cos(taz)f(z)dz = \exp(it\xi) \left\{ \cos \left( \frac{t\alpha}{\kappa} \right) + t\alpha \int_R \sin(taz) \exp \left[ - (1 - \kappa z)^{\frac{1}{\kappa}} \right] dz \right\} \mathbf{A}$$

Now integrating IInd term by parts  $\Rightarrow$

$$\int_R i \sin(taz)f(z)dz = \exp(it\xi) \left\{ \left[ i \sin \left( \frac{t\alpha}{\kappa} \right) - it\alpha \int_R \cos(taz) \exp \left[ - (1 - \kappa z)^{\frac{1}{\kappa}} \right] dz \right] \right\}$$

**B**

**A+B  $\Rightarrow$**

$$\phi_Z(t; \xi, \alpha, \kappa) = \exp(it\xi) \left\{ \exp\left(\frac{it\alpha}{\kappa}\right) - it\alpha \int_R \exp[itaz - (1 - \kappa z)^{\frac{1}{\kappa}}] dz \right\} \quad (6)$$

Adding the integral of each product obtained by multiplying each term of Taylor's series expansion of  $\exp(itaz)$  with  $\exp[-(1 - \kappa z)^{\frac{1}{\kappa}}]$  and substituting  $(1 - \kappa z)^{\frac{1}{\kappa}} = y$  in the integrals of the products leads to the CHF of type-3 EVD as:

$$\phi_X(t; \xi, \alpha, \kappa) = \exp\left[it\left(\xi + \frac{\alpha}{\kappa}\right)\right] \sum_{r=0}^{\infty} \frac{(-it\alpha)^r}{r! \kappa^r} \Gamma(1 + r\kappa); r = 0, 1, 2, \dots \quad (7)$$

**CHF of type-2 EVD:** In this work, the CHF of the Fréchet distribution is derived by the direct and lucid methodology discussed in the previous section. The CHF can be derived as:

$$\phi_X(t; \xi, \alpha, \kappa) = \int_R [\cos(tx)f(x) + i\sin(tx)f(x)] dx \quad (8)$$

Integrating 1st term by parts  $\Rightarrow$

$$\int_R \cos(tx)f(x) dx = \cos(t\xi) + t \int_R \sin(tx) \exp\left[-\left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right] dx$$

Now integrating 2nd term by parts  $\Rightarrow$

$$\int_R i\sin(tx)f(x) dx = i\sin(t\xi) - it \int_R \cos(tx) \exp\left[-\left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right] dx$$

D  
 C+D  $\Rightarrow$

$$\phi_X(t; \xi, \alpha, \kappa) = \exp(it\xi) - it \int_R \exp\left[itx - \left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right] dx \quad (9)$$

Adding the integral of each product obtained by multiplying each term of Taylor's series expansion of  $\exp(itx)$  with  $\exp\left[-\left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right]$  and substituting  $\left(\frac{x-\xi}{\alpha}\right)^{-\kappa} = y$  in the integrals of the products leads to the CHF of Fréchet distribution as:

$$\phi_X(t; \xi, \alpha, \kappa) = \exp(it\xi) \sum_{r=0}^{\infty} \frac{(it\alpha)^r}{r!} \Gamma\left(1 - \frac{r}{\kappa}\right); r = 0, 1, 2, \dots \quad (10)$$

The recent paper by Nadarajah and Pogány<sup>9</sup> obtained the CHF of type-2 EVD that uses the integral referred to as the complex parameter Kratzel function. The methodology applied by us is direct, unique and lucid and the expression for characteristic function of type-2 EVD derived here are in closed form and are simpler than that obtained by Nadarajah and Pogány<sup>9</sup> and by Muraledharan et.al<sup>10</sup> and hence suitable for further statistical applications.

**MGF of type-3 EVD:** The MGF,  $M_X(\theta; \xi, \alpha, \kappa)$  of reversed Weibull distribution ( $\kappa > 0$ ) can be computed as

$$M_Z(\theta; \xi, \alpha, \kappa) = \exp(\theta\xi) \int_R (1 - \kappa z)^{\frac{1}{\kappa}-1} \exp[\theta\alpha z - (1 - \kappa z)^{\frac{1}{\kappa}}] dz \quad (11)$$

$\theta$  – arbitrary real constant

Adding the integral of each product obtained by multiplying each term of Taylor's series expansion of  $\exp(\theta\alpha z)$  with  $(1 - \kappa z)^{\frac{1}{\kappa}-1} \exp[-(1 - \kappa z)^{\frac{1}{\kappa}}]$  and substituting  $(1 - \kappa z)^{\frac{1}{\kappa}} = y$  in the integrals of the products leads to the MGF of type-3 EVD as

$$M_X(\theta; \xi, \alpha, \kappa) = \exp(\theta\xi) \sum_{r=0}^{\infty} \frac{[\theta\alpha]^r}{r! \kappa^r} \Gamma(1 + r\kappa); r = 0, 1, 2, \dots \quad (12)$$

The raw moments of the reversed Weibull distribution can be obtained from  $M_X(\theta; \xi, \alpha, \kappa)$ . I.e.

$$M_X^{(n)}(0; \xi, \alpha, \kappa) = \mu'_n = \left[ \frac{d^n M_X(\theta; \xi, \alpha, \kappa)}{d\theta^n} \right]_{\theta=0} \quad (13)$$

$\mu'_n$  –  $n^{\text{th}}$  raw moment

**MGF of type-2 EVD:** The MGF of Fréchet distribution can also be computed by the same methodology as:

$$M_X(\theta; \xi, \alpha, \kappa) = \int_R \frac{\kappa}{\alpha} \left(\frac{x-\xi}{\alpha}\right)^{-\kappa-1} \exp\left[\theta x - \left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right] dx \quad (14)$$

Adding the integral of each product obtained by multiply each term of Taylor's series expansion of  $\exp(\theta x)$  with  $\frac{\kappa}{\alpha} \left(\frac{x-\xi}{\alpha}\right)^{-\kappa-1} \exp\left[-\left(\frac{x-\xi}{\alpha}\right)^{-\kappa}\right]$  and substituting  $\left(\frac{x-\xi}{\alpha}\right)^{-\kappa} = y$  in the integrals of the products leads to the MGF of Fréchet distribution as

$$M_X(\theta; \xi, \alpha, \kappa) = \exp(\theta\xi) \sum_{r=0}^{\infty} \frac{(\theta\alpha)^r}{r!} \Gamma\left(1 - \frac{r}{\kappa}\right); r = 0, 1, 2, \dots \quad (15)$$

## Discussion

The CHFs and MGFs of reversed Weibull ( $\kappa > 0$ ;) and Fréchet ( $\kappa < 0$ ) distributions derived here are new and are in simplest closed forms. The methodology is also unique, lucid and direct. The CHFs satisfy the tests for a function to be a characteristic function<sup>11,12</sup> such as: i. That  $\phi_X(t)$  must be continuous in  $t$ , ii. That  $\phi_X(t)$  is defined in every finite  $t$  interval, iii. 3) That  $\phi_X(0) = 1$  and 4) That  $\phi_X(t)$  and  $\phi_X(-t)$  shall be conjugate quantities.

CHFs are very useful some of which to mention are to, develop sampling distributions based on the method of characteristic functions, generate cumulants, apply in the theory of decomposability of random variables etc. The MGFs can also find applications such as in inversion, convolution of individual probability density functions of independent random variables,

generation of moments of random variables, in continuity theorem etc.

## Conclusion

The characteristic and moment generating functions of generalised extreme value distribution (GEV) for its shape parameter  $\kappa > 0$  (reversed Weibull) and  $\kappa < 0$  (Fréchet distribution) are derived for the first time by a direct, lucid and unique methodology. The expressions are new and simple. The CHFs satisfied the tests for a function to be a characteristic function and the MGFs are able to generate all the raw moments of their respective probability distributions.

## Acknowledgement

The authors are much grateful to the Centre for Marine Technology and Ocean Engineering (CENTEC), Instituto Superior Técnico, Technical University of Lisbon, Lisboa, Portugal for technical support and financial assistance.

## References

1. Jenkinson A.F. (1955). The Frequency Distribution of the Annual Maximum (or Minimum) Values of Meteorological Elements. *Quart. J. Roy. Met. Soc.*, 81(348), 158-171.
2. Jenkinson A.F. (1969). Estimation of Maximum Floods. *World Meteorological Organization, Technical Note*, 98, 183-257.
3. Natural Environment Research Council (1975). Flood Studies Report. 1, London.
4. Sinclair C.D. and Ahmad M.I. (1988). Location-Invariant Plotting Positions for PWM Estimation of the parameters of GEV Distribution. *J.Hydro.*, 99(3-4), 271-279.
5. Otten A. and Van Monfort M.A.J. (1980). Maximum-Likelihood Estimation of the generalised-extreme value distribution parameters. *J. Hydro.*, 47(1), 187-192.
6. Prescott P. and Walden A.T. (1980). Maximum Likelihood Estimation of the parameters of the Generalised Extreme-Value Distribution. *Biometrika*, 67, 723-724.
7. Prescott P. and Walden A.T. (1983). Maximum Likelihood Estimation of the parameters of the three parameter Generalised Extreme-Value Distribution from Censored Samples. *J. Stat. Comp. Simul.*, 16(3-4), 241-250.
8. Hosking J.R.M., Wallis J.R. and Wood E.F. (1985). Estimation of the Generalised Extreme Value Distribution by the Method of Probability-Weighted Moments. *Technometrics*, 27(3), 251-261.
9. Nadarajah S. and Pogány T.K. (2013). On the Characteristic Functions for Extreme Value Distributions. *Extremes*, 16, 27-38.
10. Muraleedharan G., Guedes C. and Lucas C. (2011). Characteristic and Moment Generating Functions of Generalised Extreme Value Distribution (GEV). Sea Level Rise, Coastal Engineering, Shorelines and Tides. Chapter-14, Nova Science Publishers, New York, 269-276. (ISBN: 978-1-61728-655-1)
11. Maurice George Kendall (1946). The Advanced Theory of Statistics. Distribution Theory, 1, Charles Griffin and Co. Ltd., London, 433.
12. Ochi Michel K. (1998). Ocean Waves: Ocean Technology Series-6, Cambridge University Press. Cambridge, 319.